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# The role of low-fat diets in body weight control: a meta-analysis of *ad libitum* dietary intervention studies†

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**OBJECTIVES:** Low-fat high-carbohydrate diets are recommended to prevent weight gain in normal weight subjects and reduce body weight in overweight and obese. However, their efficacy is controversial. We evaluated the efficacy of *ad libitum* low-fat diets in reducing body weight in non-diabetic individuals from the results of intervention trials. **DESIGN:** Studies were identified from a computerized search of the Medline database from January 1966 to July 1999 and other sources. Inclusion criteria were: controlled trials lasting more than 2 months comparing *ad libitum* low-fat diets as the sole intervention with a control group consuming habitual diet or a medium-fat diet *ad libitum*.

**MAIN OUTCOME MEASURES:** Differences in changes in dietary fat intake, energy intake and body weight. Weighted mean differences for continuous data and 95% confidence intervals (CIs) were calculated.

**RESULTS:** Two authors independently selected the studies meeting the inclusion criteria and extracted data from 16 trials (duration of 2–12 months) with 19 intervention groups, enrolling 1910 individuals. Fourteen were randomized. Weight loss was not the primary aim in 11 studies. Before the interventions the mean proportions of dietary energy from fat in the studies were 37.7% (95% CI, 36.9–38.5) in the low-fat groups, and 37.4% (36.4–38.4) in the control groups. The low-fat intervention produced a mean fat reduction of 10.2% (8.1–12.3). Low-fat intervention groups showed a greater weight loss than control groups (3.2 kg, 95% confidence interval 1.9–4.5 kg;  $P < 0.0001$ ), and a greater reduction in energy intake (1138 kJ/day, 95% confidence interval 564–1712 kJ/day,  $P = 0.002$ ). Having a body weight 10 kg higher than the average pre-treatment body weight was associated with a  $2.6 \pm 0.8$  kg ( $P = 0.011$ ) greater difference in weight loss.

**CONCLUSION:** A reduction in dietary fat without intentional restriction of energy intake causes weight loss, which is more substantial in heavier subjects.

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**Keywords:** low-fat diets; weight loss; obesity; humans; body weight regulation

## Introduction

Dietary guidelines recommend a reduction in total dietary fat content to less than 30% of the energy intake to help reduce the prevalence of obesity, ischaemic heart disease and certain cancers. However, the robustness of the data providing the scientific evidence for a causal link between dietary fat and obesity has been challenged. Katan *et al* have questioned the importance of low fat diets in the prevention and treatment of obesity,<sup>1</sup> and Willett has stated that 'Diets high in fat do not appear to be the primary cause of the prevalence of excess body fat in our society, and reductions in [dietary] fat will not be the solution' and 'in the longer term, fat consumption

within the range of 18–40% appears to have little if any effect on body fatness'.<sup>2</sup> These views have been disseminated to the public, and many now advocate that saturated fat be replaced by monounsaturated fat rather than by carbohydrate. If this advice is followed the current fat consumption, which already exceeds the recommended level, will be maintained or even further increased. Thus knowing whether dietary fat facilitates body weight gain and obesity, and whether reducing dietary fat can aid weight loss is crucial to formulating nutritional advice and policies.

Evidence based on observational studies looking for associations between habitual dietary fat intake and body fatness can be confounded by lack of control for factors such as physical activity, smoking etc, but can also have other limitations due to the reliance on information on dietary intakes given by the subjects under examination. Accurate information on dietary fat intake is difficult to achieve in populations that have been recommended to reduce fat intake, because they may eat a diet closer to the recommendations during the dietary survey or specifically underreport fat intake.<sup>3</sup> It is well established that overweight and

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obese subjects underreport their energy intake by 30–40%,<sup>4</sup> and fat may be over-represented in this under-reporting.<sup>5</sup> Studies in health conscious populations (eg United States) show that high-fat foods are under-reported, whereas low fat foods are overreported,<sup>6</sup> and it seems to be easier to demonstrate positive associations between dietary fat intake and subsequent weight changes in less health conscious populations (eg China<sup>7</sup>).

A large body of short-term studies on appetite and energy intake shows that fat is less satiating than carbohydrate and protein when compared Joule for Joule, and that high-fat foods are more likely than low-fat foods to induce passive overconsumption and weight gain.<sup>8–11</sup> Longer term low-fat intervention studies show large variability and the inconsistent outcomes have produced the aforementioned scepticism about the effectiveness of low-fat diets. We therefore found it pertinent to look more systematically at the evidence from *ad libitum* intervention studies to examine the effect of reducing dietary fat on body fatness. Hence, the aim of this review was to conduct a meta-analysis of controlled intervention studies of low fat *ad libitum* diets in order to assess the effects on body weight, and furthermore to elucidate factors previously reported to influence weight loss such as pre-treatment body weight and degree of dietary fat reduction.<sup>12</sup>

## Methods

### Literature search

The initial search of the literature from January 1966 to July 1999 was conducted using Medline (US National Library of Medicine), for studies having the keywords 'body weight', 'weight loss', 'weight reduction', 'body weight changes', 'fat restricted', '*ad libitum* low fat diet', 'low-fat diet', 'low-cholesterol diet' or 'reduced fat'. In addition to this search, extensive cross referencing searches were also performed and published abstracts were also evaluated for inclusion. This procedure was carried out by two independent investigators.

### Inclusion criteria

The initial search revealed 218 publications. Studies were then excluded if the duration of the intervention period was less than 2 months ( $n=47$ ), if the low-fat diet was isocaloric (ie energy intake was adjusted to maintain body weight,  $n=9$ ), if total energy intake was restricted ( $n=48$ ), if there were other interventions which could affect weight loss ( $n=12$ ), if the study involved patients with non-insulin dependent diabetes mellitus ( $n=6$ ), if drugs were given which may have affected weight loss ( $n=5$ ), if no proper control group existed ( $n=19$ ) or if no numerical values for the pre-treatment and final body weights

and/or the pre-treatment and final dietary fat intake were provided ( $n=41$ ). Where results from the same study are reported in more than one publication, the data are only included once ( $n=15$ ). A total of 16 publications reporting 19 intervention groups met the inclusion criteria. Study characteristics are shown in Table 1. These studies included a total of 1910 individuals, 62% women and 38% men.

### Study description

Three studies were carried out in women with an increased risk of breast cancer or breast dysplasia;<sup>13–15</sup> eight studies reported in seven publications assessed diets in prevention or treatment of cardiovascular disease,<sup>16–22</sup> three studies examined weight loss diets,<sup>23–25</sup> and five studies reported in three publications examined effects of diet on both blood lipids and body weight.<sup>26–28</sup> Fourteen studies were randomized controlled trials, 13 using a parallel design and one a cross-over design. The control groups were either advised to maintain their regular diet or to consume a diet with a fat content as in the background population. Data was only included from intervention periods, rather than longer-term follow-up data. We would have preferred to use body mass index (BMI, weight/height<sup>2</sup>) to get a better proxy for body fatness than provided by body weight, but BMI was not reported in all the studies and its estimation would have increased the error in variation. Mean body mass index of the studies ranged from 21 to 30 kg/m<sup>2</sup>. Data abstraction was performed by two separate investigators, and the results cross checked so as to minimize potential investigator bias in data abstraction.

### Statistical analyses

The standard deviation of the difference in weight change between the intervention and control groups was estimated by standard equations from the reported standard error, standard deviations or 95% confidence intervals (CIs) of the changes within the groups and differences between the groups. Standard methods of meta-analysis<sup>29</sup> were used to carry out a chi-square test of homogeneity of study means, a fixed effects estimate, standard error and test of the difference in weight change between intervention and control diets, and a random effects estimate, standard error and test of the same difference in weight change. A linear regression analysis was conducted to test for relationships between differences in changes in body weight in intervention and control groups and the following independent variables: average pre-treatment body weight, average age, gender indicator (percent of male subjects > 50%, y/n), duration of trial and difference between intervention and control groups in change in percentage of energy intake from fat. A quadratic effect of duration was considered to allow for the possibility that weight loss was not constant

Table 1 Characteristics of 19 controlled trials of *ad libitum* low-fat diets on body weight

Reference	Goal	Duration	Group	Gender (F/M)	Change in dietary fat energy, % (s.d.)	Weight change (kg) (s.d.)	Weight loss difference (kg) (95% CI)	Change in energy intake (MJ/day) (s.d.)
Lee-Han et al <sup>13</sup>	Breast dysplasia	12 months	I C	29/0 28/0	-10.6 (11.4) + 0.2 (9.5)	-0.9 (12.3) 0.6 (7.3)	-1.5 (-6.7-3.7)	-3.8 (9.0) + 0.9 (11.3)
Boyd et al <sup>16</sup>	Blood lipids	12 months	I C	100/0 106/0	-16.0 (10.6) -2.0 (8.5)	-1.0 (9.8) 0.0 (9.4)	-1.0 (-3.6-1.6)	-3.7 (9.2) 0 (10.5)
Buzzard et al <sup>14</sup>	Stage II breast cancer	3 months	I C	17/0 11/0	-15.6 (9.4) -3.0 (7.7)	-2.8 (2.9) -1.3 (2.7)	-1.5 (-3.6-0.6)	-8.3 (9.7) NR <sup>a</sup>
Bloemberg et al <sup>17</sup>	Blood lipids	26 weeks	I C	0/39 0/41	-5.0 (6.5) -1.5 (5.9)	-0.9 (2.7) + 0.1 (1.9)	-1.0 (-2.0-0.0)	NR <sup>a</sup> NR <sup>a</sup>
Sheppard et al <sup>23</sup>	Weight loss	12 months	I C	171/0 105/0	-17.6 (8.5) + 1.7 (7.8)	-3.0 (4.8) 0.4 (3.5)	-3.4 (-4.4--2.4)	-7.7 (7.0) -2.5 (8.1)
Baer <sup>18</sup>	Blood lipids	12 months	I C	0/33 0/37	-7.0 (24.3) -1.0 (38.0)	-5.0 (14.9) 1.0 (21.9)	-6.0 (-14.7-2.7)	NR <sup>a</sup> NR <sup>a</sup>
Hunninghake et al <sup>19</sup>	Blood lipids	9 weeks	I C	40/65	-15.6 (14.7) 0.0 (14.7)	-1.4 (36.1) 0.0 (36.1)	-1.4 (-11.4-8.6)	NR <sup>a</sup> NR <sup>a</sup>
Kasim et al <sup>20</sup>	Blood lipids	12 months	I C	34/0 38/0	-18.7 (9.1) -1.8 (10.0)	-3.4 (16.1) -0.8 (16.9)	-2.6 (-10.0-4.8)	-6.2 (12.1) -3.5 (11.7)
Raben et al <sup>21</sup>	Blood lipids	11 weeks	I C	6/18 8/16	-11.8 (12.7) 0.0 (4.8)	-1.4 (1.9) 0.0 (1.0)	-1.4 (-2.3--0.5)	NR <sup>a</sup> NR <sup>a</sup>
Pritchard et al <sup>24</sup>	Body composition	12 weeks	I C	0/18 0/19	-12.8 (2.1) + 0.7 (1.9)	-6.4 (10.8) 0.0 (12.3)	-6.4 (-13.8-1.0)	-14.2 (12.8) + 1.7 (10.3)
Siggaard et al <sup>25</sup>	Body composition	12 weeks	I C	2/47 2/14	-11.0 (10.7) -1.6 (6.6)	-4.2 (2.7) 0.8 (1.9)	-5.0 (-6.5--3.5)	-2.1 (10.0) + 0.4 (10.4)
Simon et al <sup>15</sup>	Breast cancer prevention	12 months	I C	65/0 68/0	-18.0 (8.8) -0.9 (10.0)	-1.6 (22.1) -2.3 (17.4)	+ 0.7 (-6.1-7.5)	-5.4 (12.7) -1.4 (7.8)
Weststrate et al <sup>26</sup>	Body weight blood lipids	6 months	I C	58/59 51/52	-3.0 (4.6) + 5.0 (6.5)	+ 0.4 (14.7) + 1.1 (14.7)	-0.7 (-4.6-3.2)	-1.3 (18.1) + 3.7 (16.9)
Stefanick et al <sup>22</sup>	Blood lipids females	12 months	I C	46/0 45/0	-5.7 (7.4) -0.2 (6.7)	-2.7 (3.5) + 0.8 (4.2)	-3.5 (-5.1--1.9)	-3.9 (6.2) -0.3 (6.4)
Stefanick et al <sup>22</sup>	Blood lipids males	12 months	I C	0/49 0/46	-8.0 (8.1) -0.7 (5.9)	-2.8 (3.5) + 0.5 (2.7)	-3.3 (-4.6--2.0)	-5.0 (9.5) -0.4 (8.4)

Table continued

Table 1 (continued)

Reference	Goal	Duration	Group	Gender (F/M)	Change in dietary fat energy, % (s.d.)	Weight change (kg) (s.d.)	Weight loss difference (kg) (95% CI)	Change in energy intake (MJ/day) (s.d.)
Skov et al <sup>27</sup>	Body weight (high protein)	6 months	I C	17/6 11/3	-7.2 (4.9) -2.6 (7.6)	-8.7 (4.5) +2.4 (3.5)	-11.1 (-13.8-8.2)	NR <sup>a</sup>
Skov et al <sup>27</sup>	Body weight (low protein)	6 months	I C	18/5 11/3	-9.6 (5.7) -2.6 (7.6)	-5.1 (3.9) +2.4 (3.3)	-7.5 (-9.9--5.1)	-2.9 (8.8) -3.3 (10.0)
Saris et al <sup>28</sup>	Body weight (simple carbohydrates)	6 months	I C	40/36 40/37	-10.2 (5.1) +0.8 (4.4)	-0.9 (3.6) +0.8 (4.1)	-1.7 (-2.9--0.5)	NR <sup>a</sup>
Saris et al <sup>28</sup>	Body weight (complex carbohydrates)	6 months	I C	40/43 40/37	-7.9 (4.6) -0.8 (4.4)	-1.8 (3.2) +0.8 (4.1)	-2.6 (-3.8--1.4)	-7.5 (10.0) -3.3 (10.0)
Total number of subjects					683/418 = 1101 504/305 = 809			

<sup>a</sup>NR, not reported.

over time, and the gender indicator was included in evaluating the effect of pre-treatment body weight to adjust for the effect of gender in interpreting pre-treatment weight.

The study by Hunninghake<sup>19</sup> used a crossover design but was treated in the analysis as if it was a parallel design. This gives an overestimate of the standard deviation, resulting in a lower meta-analysis and regression weight for that study. Because data were reported separately, the study by Stefanick et al<sup>22</sup> was treated as two separate studies, one for men and one for women. In the studies by Skov et al<sup>27</sup> and by Saris et al<sup>28</sup> there were two different low-fat intervention treatment groups, but both used the same control group. These studies were included as two separate studies, although the control group appears twice. Analyses were repeated omitting various combinations of these studies and no substantive changes in results were noted.

Statistical analyses were done using SAS (SAS Institute Inc., SAS/STAT Software: Changes and Enhancements through Release 6.12, Cary NC: SAS, 1997), with PROC MIXED in SAS used for the random effects analysis. Graphs were made using S-plus (S-plus User's Guide, Data Analysis Products Division, MathSoft, Seattle, WA, USA).

## Results

Before the interventions the mean proportions of dietary energy from fat in the studies were 37.7% (95% CI, 36.9-38.5) in the low-fat groups, and 37.4% (36.4-38.4) in the control groups. The low-fat intervention produced a mean fat reduction of 10.8% of energy (range: 3.5-19.3%), while no change occurred in the control groups. In the 12 (of 19) studies with reported energy intake data the intervention arms produced a mean change in total energy intake greater than that of the control group ranging from -3800 to +100 kJ/day and the meta-analysis showed that the low-fat interventions as compared with the control groups produced an estimated reduction in energy intake of 959 kJ/day more (95% CI, 729-1189 kJ/day;  $P < 0.0001$ ) in the fixed effects analysis and 1138 kJ/day more (564-1712 kJ/day;  $P = 0.002$ ) in the random effects analysis (Figure 1). The chi-square test of homogeneity of study means gave  $Q = 21.0$  (d.f. = 11,  $P = 0.033$ ), which indicated greater differences among study means than could be explained by random sampling, and in this case the random effects analysis is to be preferred. However, it is recommended that both the fixed effects and random effects analyses be carried out as a sensitivity check, and in this case the results are quite similar.

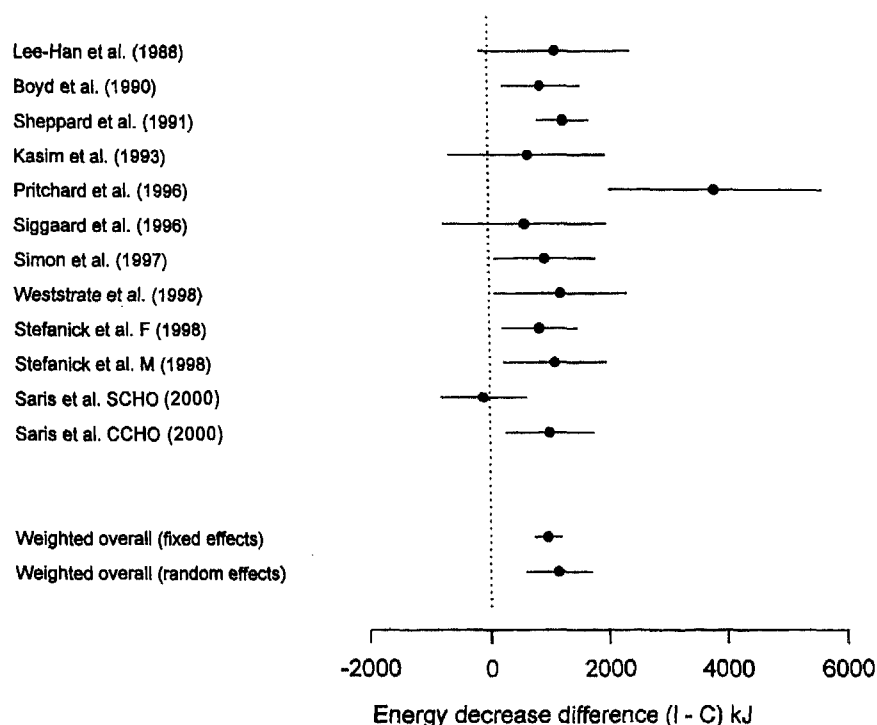
In the 19 studies included in this meta-analysis the low-fat interventions produced a weighted mean reduction in percentage of energy from fat of 10.2% (95% CI, 8.1-12.3) as compared to the control

groups. The corresponding difference in mean weight losses ranged from  $-0.7$  to  $11.1$  kg. The meta-analysis revealed that the low-fat interventions as compared with the control groups produced an estimated mean weight loss of  $2.6$  kg more (95% CI,  $2.3$ – $3.0$  kg;  $P < 0.0001$ ) in the fixed effects analysis and  $3.2$  kg more ( $1.9$ – $4.5$  kg;  $P < 0.0001$ ) in the random effects analysis. The test of homogeneity of study means gave  $Q = 93.4$  (d.f. = 18,  $P < 0.0001$ ), again indicating heterogeneity of study means and a preference for the random effects analysis. Figure 2 shows these weight losses and combined estimates along with 95% CIs. In unweighted univariate regression analyses the weight loss difference was associated only with pre-treatment average body weight adjusted for gender ( $P = 0.008$ ), but not significantly with reduction in percentage dietary fat intake ( $P = 0.08$ ), energy intake, gender or duration of trial. In weighted univariate regressions, none of the independent variables were significantly associated with weight loss. In unweighted multivariate regression both pre-treatment body weight ( $P = 0.011$ ) and gender ( $P = 0.041$ ) were significantly associated with weight loss, but these determinants were not significant in a weighted multivariate regression analysis ( $P = 0.15$ ,  $P = 0.22$ ). Excluding the high-protein arm of the Skov *et al* study<sup>27</sup> and the high-simple sugar arm of the Saris *et al*<sup>28</sup> study did not essentially change the results. The significant effect of pre-treatment body weight (adjusted for gender) in the unweighted regressions can be quantified using the coefficients from the regressions. Mean difference in

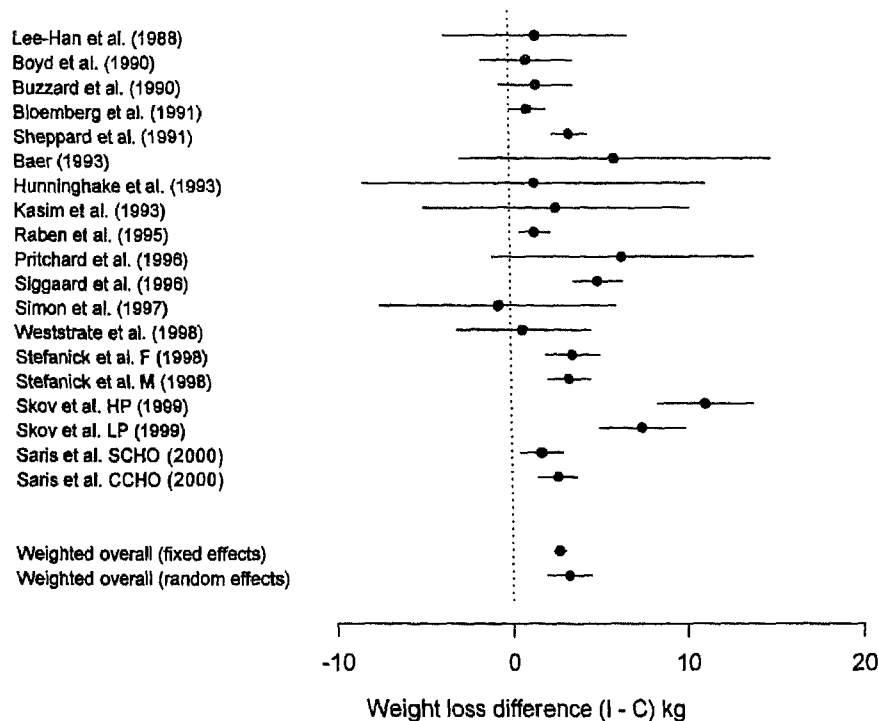
weight loss (I–C) was greater in studies with higher mean pre-treatment body weight. A  $10$  kg higher pre-treatment mean body weight was associated with a greater difference in weight loss (I–C) of about  $2.0 \pm 0.7$  kg ( $P = 0.008$ ) from the univariate regression and about  $2.6 \pm 0.8$  kg ( $P = 0.011$ ) from the multivariate regression.

## Discussion

The meta-analysis of the results from the 19 controlled *ad libitum* low-fat 2–12 months trials on weight change involving 1910 individuals of both genders shows that a reduction in the dietary energy from fat is associated with a spontaneous weight loss of  $3.2$  kg more ( $1.9$ – $4.5$  kg;  $P < 0.0001$ ) in the intervention than in the control group. The weight loss was not related to the duration of the low-fat intervention, probably because weight loss plateaued after 3–6 months. For this reason the results cannot be compared with the estimate by Bray and Popkin,<sup>11</sup> who also included trials of shorter duration and without control groups in their review. Based on 28 trials they found that a reduction of 10% in the proportion of energy from fat was associated with a reduction in weight of  $16$  g/day,<sup>11</sup> which corresponds to  $2.88$  kg over 6 months, which is within the confidence intervals of the estimate of our meta-analysis (the mean difference in percentage dietary fat reduction was 10.8% for our studies). Another meta-analysis tested the effect of the National Cholesterol Education



**Figure 1** Differences in energy intake (change in intervention minus change in control, kJ/day) with 95% confidence intervals for 12 studies included in the meta-analysis. The weighted estimates and 95% confidence intervals from the fixed effects and random effects meta-analysis are also shown.



**Figure 2** Differences in weight loss (change in intervention minus change in control; kg) with 95% confidence intervals for 19 studies included in the meta-analysis. The weighted estimates and 95% confidence intervals from the fixed effects and random effects meta-analysis are also shown.

Program's Step I and Step II low-fat diets for reduction of cardiovascular risk factors, but had more liberal inclusion criteria than the present analysis and was therefore based on 37 dietary intervention studies.<sup>30</sup> The weight loss in the intervention groups was 2.79 kg larger than in the control group, and there was a highly significant relation between reduction in dietary fat and weight loss, so that for every 1% decrease in energy from fat there was a 0.28 kg decrease in body weight.<sup>30</sup> In contrast, our meta-analysis failed to detect an association between the reduction in dietary fat and weight loss, but this may be due to the reduced power because we included fewer studies.

Another important finding was that the weight loss was dependent on the pre-treatment body weight so that each 10 kg of initial pre-treatment mean body weight related to an additional 2.0–2.6 kg weight loss compared with control groups. Bray and Popkin in their weighted regression also found a large effect of initial weight and an added effect of being male.<sup>11</sup> In individual studies it has previously been described that weight loss is positively correlated with pre-treatment body weight, such as in that of Schaefer *et al* ( $r = 0.68$ ,  $P < 0.001$ ).<sup>31</sup> Furthermore, Siggaard *et al* stratified subjects participating in a low fat *ad libitum* trial into normal weight ( $\text{BMI} < 25 \text{ kg/m}^2$ ) and overweight subjects ( $\text{BMI} > 25 \text{ kg/m}^2$ ), and found that after 12 weeks the normal weight group had lost significantly less than the overweight group (2.5 vs 5.2 kg).<sup>25</sup> Raben *et al* also reported that initial body fat was an important determinant of weight loss achieved during a low-fat diet.<sup>21</sup>

No trials with groups of subjects with a BMI of more than 30 kg/m<sup>2</sup> fulfilled our inclusion criteria, so

we cannot draw any conclusions on the potential of *ad libitum* low-fat diets to produce weight loss in obese subjects. The lack of studies in obese subjects fulfilling our inclusion criteria is somewhat surprising, and cannot explain the origins of the general view that low-fat diets are not effective in producing a clinically relevant weight loss in obese subjects.<sup>32</sup> This assumption derives from trials conducted in overweight individuals, uncontrolled trials in obese subjects and two randomized studies comparing the *ad libitum* low-fat diet with an energy restricted diet.<sup>32</sup> Shah *et al* found that the low-fat diet reduced body weight by 4.4 kg and the low-energy diet by 3.8 kg after 6 months.<sup>33</sup> Jeffery *et al* found a 6 month weight loss of 4.6 kg on the low-fat diet and 3.7 kg on the low-energy diet, and similar regain in both groups after 12 months.<sup>34</sup> These observations together with our analysis suggest that a mean weight loss of 4–5 kg can be obtained by obese subjects with a BMI around 30 kg/m<sup>2</sup>. Although it may be a modest outcome in obese subjects whose weight exceeds the normal by 35–50 kg, a 5% weight loss is considered to have an important beneficial effect on risk of morbidity.<sup>35</sup>

The reason for the larger weight loss with larger body fatness is not explained by the present study but a number of studies have suggested that genetically predisposed subjects may be more likely to gain weight on high-fat diets.<sup>36</sup> Other studies have suggested this susceptibility to be linked to a lower ability to oxidize fat.<sup>37,38</sup> If a higher susceptibility to weight gain on high-fat diets plays a role in the development of overweight and obesity, it is plausible that a

reduction in dietary fat may also produce larger weight loss in susceptible individuals. The results of this meta-analysis may also have relevance for weight maintenance in obese subjects who have lost weight on a low-energy diet. A randomized intervention study demonstrated that 2 y after a major weight loss a group of obese subjects who consumed an *ad lib* low-fat diet had regained, on average, 5.9 kg less than a group who cut down on all calories equally.<sup>39</sup> Moreover, it has been shown that persons successful at long-term weight loss and maintenance are those who continue to consume a low-energy, low-fat diet.<sup>40</sup>

Apart from the pre-treatment body weight, duration and reduction of the dietary fat content a number of other factors may have influenced the magnitude of weight loss. Firstly, it is not possible to conduct long-term randomized controlled low-fat trials in a design blinded to the participants. In studies conducted for weight loss purposes the open design may tend to overestimate the intrinsic effect of the low-fat diet because the subjects in an effort to lose weight may voluntarily restrict their total energy intake or make other efforts to lose weight. However, in 11 of the 19 studies in this meta-analysis the goal of the study was not related to weight loss, and in some of the studies weight loss was even regarded as an undesirable effect<sup>41</sup> which was tolerated only during the first 11 weeks of the study.<sup>21</sup> Consequently, it is not likely that voluntary energy restriction has confounded the outcome of the analysis severely. Secondly, in some of the studies the control groups did not receive counselling at similar intervals as the intervention groups, and it is possible that this may have contributed to the observed weight loss difference. However, in the studies by Weststrate *et al* and Saris *et al* the normal-fat control groups received the similar degree of counselling and all foods free of charge, and the outcome of these studies is within the confidence intervals of the meta-analysis. Thirdly, as with other lifestyle changes, adherence to the diet is crucial if an effect is to be observed. In the studies contributing to this meta-analysis, dietary compliance was assessed by dietary records, which allowed a calculation of energy and macronutrient intake during the trial. However, one should keep in mind that the information about the diet intake is self-reported by the participants and the validity is therefore questionable and the information may be idealized. If dietary fat intake is underreported the impact of low-fat diets on weight loss is underestimated and for a given dietary fat reduction the weight loss would be greater than estimated from the present analysis. Low-fat diets should not be blamed for lack of efficacy if the real reason for the lack of observed effect is the subjects' resumption of their former high-fat dietary habits. However, it remains to be seen whether the modest long-term effect of low-fat diets on body weight is due to non-compliance, or whether the subjects compensate when they become adapted to the low-fat diet.

In conclusion, this meta-analysis, based on 19 controlled, *ad libitum*, low-fat, 2–12 month intervention studies, shows that *ad libitum*, low-fat diets cause weight loss. The effect is more pronounced in subjects with a higher initial body weight. The meta-analysis revealed a 3.2 kg greater weight loss as a result of consuming a low-fat *ad libitum* diet. A decrease in the body weight of this magnitude in the general population could reduce the prevalence of obesity from 25% to 15%.<sup>35</sup> Even a few kilograms weight loss on a low-fat diet may, therefore, have an important public health impact, but increases in daily physical activity would also be desirable to strengthen the effect of preventive efforts.

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